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Going where traditional markers have not gone before: utility of and promise for RAD sequencing in marine invertebrate phylogeography and population genomics

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Abstract

Characterization of large numbers of single-nucleotide polymorphisms (SNPs) throughout a genome has the power to refine the understanding of population demographic history and to identify genomic regions under selection in natural populations. To this end, population genomic approaches that harness the power of next-generation sequencing to understand the ecology and evolution of marine invertebrates represent a boon to test long-standing questions in marine biology and conservation. We employed restriction-site-associated DNA sequencing (RAD-seq) to identify SNPs in natural populations of the sea anemone Nematostella vectensis, an emerging cnidarian model with a broad geographic range in estuarine habitats in North and South America, and portions of England. We identified hundreds of SNPcontaining tags in thousands of RAD loci from 30 barcoded individuals inhabiting four locations from Nova Scotia to South Carolina. Population genomic analyses using high-confidence SNPs resulted in a highly-resolved phylogeography, a result not achieved in previous studies using traditional markers. Plots of locus-specific F_{ST} against heterozygosity suggest that a majority of polymorphic sites are neutral, with a smaller proportion suggesting evidence for balancing selection. Loci inferred to be under balancing selection were mapped to the genome, where 90% were located in gene bodies, indicating potential targets of selection. The results from analyses with and without a reference genome supported similar conclusions, further highlighting RAD-seq as a method that can be efficiently applied to species lacking existing genomic resources. We discuss the utility of RAD-seq approaches in burgeoning Nematostella research as well as in other cnidarian species, particularly corals and jellyfishes, to determine phylogeographic relationships of populations and identify regions of the genome undergoing selection.

Keywords: balancing selection, estuarine, genome, Nematostella, next-generation sequencing, phylogeography

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Introduction

Population genomic approaches offer revolutionary opportunities over traditional population genetic markers to characterize species and population histories, and the genetic mechanisms of adaptation by analysing

polymorphic markers dispersed throughout the entire genome (Luikart et al. 2003; Nadeau & Jiggins 2010). Historically, methods to identify large numbers of genetic markers and characterize their geographic distribution in natural populations were labour-intensive and cost-prohibitive for almost any species, particularly those lacking extensive sequence resources. However, advances in sequencing technology in recent years have opened new avenues for the generation of large numbers of molecular markers in a panel of individuals to better characterize the ecology and evolution of traditionally nonmodel species (Rowe et al. 2011). One of these methods is restriction-site-associated DNA sequencing (RAD-seq), which combines enzymatic fragmentation of the genome with high-throughput sequencing for the generation of large numbers of SNP markers (Baird et al. 2008).

Knowing the proportion of genetic exchange among populations and the spatial distribution of genetic diversity for particular species within aquatic ecosystems is critical in order to understand biodiversity and inform conservation and management decisions (Palumbi 2003, 2004; Botsford et al. 2009). Marine and estuarine habitats are relatively poorly characterized ecosystems for which we know little about the population genetics of most resident species when compared to terrestrial systems. Current data support a spectrum of expectations for species' dispersal, and the resulting population connectivity, from nearly open to a higher degree of population genetic structure over unexpectedly small geographic distances due to local recruitment (Hauser & Carvalho 2008; Ciannelli et al. 2010). Previous expectations for connectivity relied on the pelagic larval duration (PLD) to hypothesize relative dispersal distances and thus the probability of gene flow in natural populations (Cowen et al. 2000; Bay et al. 2006; Cowen & Sponaugle 2009). However, recent studies have convincingly shown that PLD is at best weakly correlated with population genetic structure (Bradbury et al. 2008; Weersing & Toonen 2009), which may be driven by errors and uncertainties when calculating F_{ST} (Faurby & Barber 2012), making confident, accurate predictions about population connectivity in the marine environment difficult.

The attributes of genetic markers for making population genetic inferences can have substantial impacts on which hypotheses can be adequately tested. Previous reviews have discussed the relative merits and limitations of the diverse set of molecular markers available for studying population processes (Parker *et al.* 1998; Sunnucks 2000; Mariette *et al.* 2002; Brumfield 2003; Brito & Edwards 2008; Diniz-Filho *et al.* 2008). To date, a large majority of studies that characterize the population genetics of marine or estuarine species have

utilized allozymes, anonymous markers (e.g. AFLPs, RFLPs), a small number of microsatellites or a handful of sequence-based markers (e.g. mitochondrial DNA and nuclear ribosomal DNA). These markers have trade-offs that frequently balance diversity (e.g. microsatellites, AFLPs) with the ease of interpretation and ability to compare among species (e.g. sequence markers). More recent surveys discussing the utility of genetic markers have emphasized the significant advantages of single-nucleotide polymorphisms (SNPs) for population genetic studies (Brumfield 2003; Morin et al. 2004; Brito & Edwards 2008). Although SNPs have the limitation of lower diversity due to only four possible allelic states and a low mutation rate, they have clear advantages for accommodating diverse assumptions of linkage or independence of markers depending on the discovery strategy, explicit models of evolutionary change and potential for roles in functional evolution (e.g. polymorphisms in coding or promoter regions). In addition, SNPs can be readily compared among genomes (nuclear, mitochondrial, chloroplast) to utilize the underlying mutational scales to characterize evolutionary processes (Morin et al. 2004; Petit et al. 2004).

Nematostella vectensis is an anthozoan cnidarian (Cnidaria, Anthozoa, Hexacorallia, Actiniaria) common to tidally restricted pools in high marsh environments (Hand & Uhlinger 1994). In recent years, N. vectensis has emerged as a model cnidarian in molecular biology and comparative genomics due to ease of laboratory culture and the publication of its genome (Putnam et al. 2007), the first for a cnidarian. Sexual reproduction and developmental stages have been well characterized in a laboratory cultures (Reitzel et al. 2007). Eggs of N. vectensis are released in a gelatinous mass by female anemones and then externally fertilized by males. Development progresses from a fertilized egg to an early embryo within the egg mass. Subsequently, early larvae swim from the degraded egg jelly, develop into an elongated late larval stage and then settle as a fourtentacle juvenile stage within 7 days. This species holds great promise as a useful model for understanding the ecological genomics of coastal species (Darling et al. 2005) given that it is found in high marsh estuaries that are impacted by human encroachment, has been repeatedly introduced to non-native locations and has a broad geographic range likely resulting in local adaptation. N. vectensis has been collected in salt marshes along the Pacific and Atlantic coast of North America, a portion of England (Hand & Uhlinger 1994; Reitzel et al. 2008) and Brazil (Silva et al. 2010). Previous research on the population genetic structure of N. vectensis, using RAPDs, AFLPs and microsatellites, has identified significant genetic differences among major coastline regions, estuaries within each region and even among subpopulations within a single estuary (Pearson et al. 2002; Darling et al. 2004; Reitzel et al. 2008; Darling et al. 2009). These studies have also shown high variation in the relative contribution of clonal reproduction to resident populations throughout its range (Darling et al. 2009). In addition, available data suggest that N. vectensis has been introduced from the west Atlantic to the west coast of North America and England, where it receives protective status under the Wildlife and Countryside Act. Despite these insights, we lack an understanding of the phylogeography of populations in the native range due to the low resolution provided by these traditional markers. High-resolution data are critical for testing hypotheses about the historical distribution of this species, the connectivity of current populations and the source locations for introduced populations in non-native habitats. Moreover, there are currently few data to test for potential genetic adaptation in natural populations that span its large geographic range. Two previous studies (Sullivan et al. 2009; Reitzel et al. 2010) utilized expressed sequence tags generated during the sequencing of the N. vectensis genome to document polymorphisms in coding regions, particularly nonsynonymous substitutions in conserved protein domains. Their findings suggest that SNPs are present in genes that could exert a large influence on protein function. More recent work has identified substantial phenotypic variation in natural populations (Reitzel et al. in revision) highlighting the need for high-density genomic markers to provide the tools for linking genetic and phenotypic diversity in populations occupying environmental gradients that may result in phenotypic clines.

Our understanding of the genetic diversity and coarse population-level relationships for N. vectensis is representative of the general population genetic data for other cnidarians. Within the marine environment, cnidarians represent a critical taxonomic group of benthic and pelagic species for both ecological function and conservation management. The phylum Cnidaria contains corals that are ecosystem engineers and support a rich biodiversity in shallow and deep marine habitats (Jones et al. 1994; Roberts et al. 2006), but are frequently threatened by anthropogenic activities. Moreover, jellyfish have emerged as common nuisance species where population blooms dramatically impact fisheries and pelagic biodiversity (Purchell et al. 2007). For species of conservation concern, resolving genetic diversity and its structure is critical to understand the impact of human activities as well as the opportunity for recovery after disturbances (Palumbi 2003; Baums 2008). In addition, understanding genetic diversity will assist in assessing the opportunity for adaptation of populations to changing environments (Hughes et al. 2003) and, with the availability of a genome, identification of genomic

regions under selection. For groups, that are exerting negative impacts on marine communities and human economies, such as jellyfishes, high-resolution characterization of genetic diversity would markedly improve our understanding of the impacts derived from the introduction of these species to non-native areas and the composition of blooms that develop in particular locations. Despite the clear need for data to understand phylogeography and the particular regions of the genome undergoing selection, population genetic studies of cnidarians often are unable to resolve many of these questions due to the availability of only few allele-based markers (e.g. microsatellites), with the exception of a small number of species, as well as the near absence of variable sequence-based markers (Shearer et al. 2002; Bilewitch & Degnan 2011). Thus, the development and application of next-generation sequencing to the population genetics of cnidarians will bridge these critical gaps. In this respect, N. vectensis is an ideal cnidarian model in which to assess how RAD-seq, or similar genomic methods, can be utilized to characterize phylogeographic relationships among populations as well as regions of the genome under selection.

In this study, we utilized RAD-seq to characterize the genetic diversity and population genetic structure of N. vectensis individuals collected along the Atlantic coast of North America. We compared our results with and without the use of the available reference genome to assess the potential impacts of utilizing RAD-seq in nonmodel species with limited genomic data. Finally, we mapped the SNPs inferred to be under selection to the reference genome in order to identify genes or genomic regions that are likely under selection, and then, we grouped them based on potential biological function. Together, our study provides one of the first applications of RAD-seq to a marine invertebrate (see De Wit & Palumbi 2012) and highlights the utility of a reference genome in generating hypotheses for linking population and functional genomics.

Methods

Collection

Adults of N. vectensis were collected from three estuaries along the Atlantic coast of North America (Peggy's Cove, Nova Scotia; Sippewissett, Massachusetts; and Baruch, South Carolina; see Reitzel et al. (2008) for details). Briefly, individuals were sieved from loose sediments, transferred to 13% (parts per thousand) artificial seawater and transported to the laboratory. Individuals were maintained under a standard culturing protocol for N. vectensis (13%, artificial seawater, fed 2-3 times per week with freshly hatched Artemia sp.). Individuals from a common laboratory stock maintained in the Martindale laboratory (Kewalo Marine Laboratory, University of Hawaii) were originally collected from Rhode River, Maryland. In addition, this laboratory culture served as the source population from which the *N. vectensis* genome was sequenced. When necessary, individual clonal lines were generated by transverse bisection to yield adequate amounts of genomic DNA.

Molecular laboratory methods

Individual anemones or pooled individuals developed through bisected clonal lines were starved for at least 3 days prior to genomic DNA extraction to minimize potential contamination from food sources. Genomic DNA for nine individuals from each of the Nova Scotia and Massachusetts populations, and six from the Maryland and South Carolina populations, was extracted with the Qiagen DNAeasy kit (Qiagen). Genomic DNA quality was checked by visual inspection on an agarose gel and with a ND-1000 Nanodrop spectrophotometer (Nanodrop Technologies). DNA concentration was also determined with a Nanodrop spectrophotometer. Ten micrograms of high-quality (260/280 > 1.8) genomic DNA per individual was submitted to Floragenex Inc. for library preparation and sequencing. Individual libraries were produced from DNA digested with a high-fidelity SbfI restriction enzyme and barcoded with 5-base pair sequence tags. Libraries were sequenced on a single lane of an Illumina GAIIX sequencer.

Data QC & QA and SNP calling

Sequencing data were filtered using the program PRINSEQ v0.18 (Schmieder & Edwards 2011). All sequence reads (i.e. individual fragments of contiguous nucleotide bases) were trimmed to a length of 31 bp; shorter reads were discarded. Reads with ambiguous characters or with mean Phred quality score (Ewing & Green 1998; Ewing *et al.* 1998) lower than 20 (base call accuracy lower than 99%) were also discarded (Huse *et al.* 2007).

Reads were aligned to the reference genome of *N. vectensis* (v1.0, http://genome.jgi-psf.org/Nemve1/Nemve1.home.html) using BOWTIE v0.12.7 (Langmead *et al.* 2009). Only reads that produced a unique best alignment to the genome (in terms of the smallest number of mismatches and the highest Phred score of the mismatch positions), with at most 3 mismatches, were retained. Aligned reads were processed in the program STACKS v0.998 (Catchen *et al.* 2011)—a tool used to form stacks of identical unique sequences from each

individual, identify loci by aligning homologous stacks, generate genotypes and match loci among individuals. High-confidence SNP calls in STACKS are performed using a maximum-likelihood framework that accounts for sources of error inherent to RAD markers (i.e. sequencing error, variable depth of coverage) (Hohenlohe et al. 2010; Catchen et al. 2011). A minimum depth of four reads per stack (i.e. eight per locus) was enforced. Significantly high-repetitive stacks were discarded by implementing the deleveraging algorithm, as these likely represent sequencing errors, duplications or repetitive regions. The deleveraging algorithm assumes similar depths for stacks originating from a common locus (Catchen et al. 2011). No mismatches among loci were allowed when creating the catalog of all the loci identified among the sampled individuals. In a similar manner, reads were processed without the use of a reference genome in order to evaluate the effects of the lack of this resource in downstream analysis. A maximum number of two mismatches was allowed among loci within each individual. Loci with more than two alleles per SNP per individual were discarded as these are considered methodological artefacts in diploid organisms or products from multiple-copy elements in the genome. Hereafter, we refer to the loci identified in this analysis as RAD markers.

The reference genome of N. vectensis (Putnam et al. 2007) was sequenced from the offspring of two parent strains originally collected from Rhode River, Maryland, USA, which is one of the populations sampled for this study. The use of this reference genome to process sequence reads by retaining only those that produce unique alignments to it could introduce a form of ascertainment bias (i.e. markers present in individuals from the Maryland population being more likely to be included in the analyses than others). To assess the effect of this potential source of bias, we tested for significant differences in the average number of reads with one reported alignment to the genome and the number of RAD markers, per individual, among populations. To account for the variability in number of reads among individuals, we randomly resampled the sets of reads in each individual in order to normalize them to the set with the smallest number of reads (56 851), using the PERL-script DAISYCHOPPER v0.6 (available from www.genomics.ceh.ac.uk/GeneSwytch).

Clone detection

Due to the capability of *N. vectensis* to reproduce asexually, we tested for the presence of clones in our data set by comparing the percentage of genotypic distances among individuals within each population. To account

for the possibility that the observed differences were caused by variability in sequencing coverage of particular markers among individuals and/or SNP calling errors, we established an arbitrary cut-off value of 95% for the percentage of genotypic pairwise distances (i.e. individuals with genotypic distances smaller than 5% are considered potential clones). This is a conservative threshold considering that the probability of a given genotype for any individual in our study was calculated to be less than 1×10^{-9} (Arnaud-Haond & Belkhir 2007; Arnaud-Haond *et al.* 2007). To evaluate the effect of the presence of potential clones in the data set, all subsequent analyses were performed comparatively using genome-aligned or unaligned reads and with or without potential clone individuals.

Detection of markers under selection

To identify potential markers in genomic regions subject to selection, we used the F_{ST} outlier method (Beaumont & Nichols 1996) implemented in the program LOSITAN (Antao et al. 2008). This method utilizes the observed allele frequencies of SNPs to estimate expected heterozygosities and global unbiased F_{ST} values (Weir & Cockerham 1984; Cockerham & Weir 1993) to simulate an expected neutral distribution for F_{ST} , assuming an island model of migration (Wright 1931). One million simulations were performed assuming an infinite alleles mutation model. 95% confidence intervals were built around the simulated mean neutral F_{ST} . SNPs with F_{ST} values significantly greater than expected under neutrality were considered candidates for positive selection. Conversely, SNPs with F_{ST} values significantly smaller than expected under neutrality were considered candidates for balancing selection (Beaumont & Nichols 1996). RAD markers containing SNPs with conflicting selection classifications (e.g. one SNP candidate neutral and another candidate balancing, in the same marker) were excluded from the analyses to avoid ambiguities.

Candidate markers under selection

Candidate markers under balancing selection that were common among all four analyses (genome-aligned or unaligned reads, with or without potential clones) were mapped to the reference genome of *N. vectensis* (Putnam *et al.* 2007). Position of each marker was annotated whether it was located in an annotated gene body (intron or exon) or close to the nearest annotated gene in the current version of the genome. When the marker was located in an intergenic region, we identified the closest gene and quantified the distance to this gene. Selected genes were then tentatively assigned a name based on U.S. Department of Energy Joint Genome

Institute (JGI) annotations or on sequence similarity to available protein sequences assessed through BLASTp searches at the U.S. National Center for Biotechnology Information (NCBI) REFSEQ. Gene Ontology (GO) and Kyoto Encyclopedia of Genes and Genomes (KEGG) pathway assignments for each selected protein were identified using the program Blast2GO v2.5.1 (Conesa et al. 2005). The results from GO analysis were grouped based on 'biological process' to cluster potential shared functions for these proteins.

Demographic inferences

Inferences of demographic statistics were carried out using candidate neutral markers only. Only one SNP per RAD marker was taken into account to avoid violating the assumption of independence among markers. Only biallelic SNPs were included in order to simplify the calculations and fit the assumptions of the software utilized for the analyses. As indicated above, all inferences were performed comparatively using genome-aligned or unaligned reads and with or without potential clone individuals.

To evaluate the validity of putative populations defined by their sampling location, we inferred population structuring through a principal component analysis (PCA) using the software EIGENSOFT v4.2 (Patterson et al. 2006; Price et al. 2006). We evaluated the significance of the identified principal components through Tracy-Widom statistics (Tracy & Widom 1994; Johnstone 2001). The statistical significance of the differences between identified populations was evaluated via a chi-square test. The summing of ANOVA statistics of genetic differentiation between pairs of populations along each eigenvector approximates a chi-square distribution with degrees of freedom equal to the number of eigenvectors (Patterson et al. 2006; see EIGENSOFT documentation). We also inferred population structuring (historical lineages) by maximizing the posterior probability of the genotypic data, given a set number of clusters (K). This method is known as Bayesian population clustering and is implemented in the program STRUCTURE v2.3.2 (Pritchard et al. 2000; Falush et al. 2003) available in the Bioportal (Kumar et al. 2009). The admixture model was used with uncorrelated allele frequencies. The MCMC was run for 1 100 000 repetitions (burnin period 1 000 000). Values for K were evaluated from 1 to 5 (10 replicates each). The optimal value of K was selected using the program STRUCTURE HARVESTER v0.6.92 (Earl & Vonholdt 2012) according to the ad hoc ΔK statistic (Evanno et al. 2005), which is the second-order rate of change of the likelihood function. STRUCTURE results were visualized using the program DISTRUCT v1.1 (Rosenberg 2004).

The overall RAD marker variability was compared among individuals within each population. A quantitative measure of this variation was obtained by estimating four commonly used genetic diversity indexes: the proportion of polymorphic SNPs, the mean observed heterozygosity, the mean expected heterozygosity and the mean number of alleles. These indexes were calculated with the R-package POPGENKIT v1.0 (Rioux Paquette 2011).

Genetic differentiation among populations was measured using the unbiased $F_{\rm ST}$ estimator $\hat{\theta}$ (Weir & Cockerham 1984) (here referred to as $F_{\rm ST}^{\rm W.C.}$) and the asymptotically consistent estimator \hat{F} (Reich et~al.~2009) (here referred to as $F_{\rm ST}^{\rm R.}$) using custom scripts in R. \hat{F} has been shown to consistently yield accurate estimates of population differentiation at small sample sizes (n < 6) when large numbers of loci (>100) are available (Willing et~al.~2012). A correction that accounts for potential inbreeding effects on \hat{F} (Reich et~al.~2009) (here referred to as $F_{\rm ST}^{\rm Rcor}$), which could be prevalent in N.~vect-ensis due to possible small effective population sizes, was also applied. Confidence intervals were calculated for each estimator based on 1000 bootstrap replicates.

In order to generate useful sequence matrices for phylogeographic analyses, the nucleotide identity data from individual homozygous SNP loci (variable among individuals) were sorted and concatenated following the procedures suggested by Emerson *et al.* (2010). Phylogenetic inferences of evolutionary relationships were performed through the implementation of statistical methods following the maximum-likelihood criterion as implemented in PHYML v3.0 (Guindon *et al.* 2010). The general time-reversible model (GTR) of nucleotide substitution (Tavare 1986) was assumed. Topological robustness was assessed through 1000 nonparametric bootstrap replicates. Trees were visualized and edited in the program FigTree v1.3.1 (Rambaut 2009).

Results

SNP discovery and clone detection

The mean number of sequence reads obtained per individual was 160 409 (95% CI \pm 31 084; SD = 85 127; n = 30), and individual values ranged between 56 851 and 353 084 reads. On average, 1721 reads (1%; 95% CI \pm 31 084; SD = 85 127; n = 30) were discarded as low quality (Fig. S1, Supporting information). An average of 114 071 reads (95% CI \pm 22 240; SD = 60 907; n = 30) had a unique alignment to the reference genome, representing ~71% of all reads (Fig. S1, Supporting information). Approximately 4% of the reads failed to produce an alignment, and 24% were discarded due to having more than one reportable alignment. To avoid any possible downstream analytical biases on the accuracy of

population statistics estimates and their uncertainty, for example Hinrichs & Suarez (2005), only markers present in all individuals were retained. Reads processed without alignment to the reference genome yielded 20% more RAD markers than the genome-aligned reads (see Table 1 and Fig. S2, Supporting information for additional details). The percentage increase in the number of polymorphic RAD markers and the number of SNPs per individual was 88% and 107%, respectively. However, there was an overall slight reduction in the number of polymorphic RAD markers (14% less) and the number of SNPs (18% less) that were shared among all individuals in this unaligned analysis.

The ascertainment bias analysis performed to address the possible effect of using the reference genome to process the sequence reads showed that, when comparing among populations, individuals from Maryland (same population as the source of the reference genome) had the largest number of retained reads. However, there were no significant differences in the average number of identified RAD markers between the Maryland and Massachusetts populations, and only marginal differences between these and the populations from Nova Scotia or South Carolina ($\alpha = 0.05$, see 95% confidence intervals in Fig. S3, Supporting information).

There were eight individuals identified as potential clones in three populations: two in Massachusetts, five in Nova Scotia and one in Maryland. Not a single potential clone pair shared identical genotypes. The percentage of pairwise genotypic similarities among potential clones ranged between 99.0% and 99.9% (mean = 99.5%, 95% CI \pm 0.1; SD = 0.3; n = 13). In contrast, the genotypic distances among nonpotential clones ranged between 61.2% and 86.5% (mean = 73.3%, 95% CI \pm 1.5; SD = 7.3; n = 89). As mentioned in the Methods section, all analyses were also performed after excluding these potential clone individuals from the data set. The results of these analyses were almost identical to the ones obtained when all sampled individuals were included (including potential clones). When the potential clones were excluded, the sequence reads processed without alignment to the reference genome yielded 17% more RAD markers than the genome-aligned reads (see Table 1). This produced a percentage increase in the number of polymorphic RAD markers and the number of SNPs, per individual, of 77% and 99%, respectively. However, the number of polymorphic RAD markers and the number of SNPs shared among all individuals remained virtually unaltered (changes were less than 4%).

Detection of markers under selection

Overall, there were approximately 200 candidate neutral markers and approximately 70 candidate balancing

Table 1 RAD marker statistics per analyses with or without clones, and using genome-aligned or unaligned reads

	Potential clones included ($n = 30$) Genome Aligned N	30) No Genome Aligned	Potential clones removed ($n = 22$) Genome Aligned N	22) No Genome Aligned
Mean number of RAD markers per individual Mean depth of coverage per RAD marker per individual	2305 (95% CI \pm 71; SD = 193) 48× (95% CI \pm 8; SD = 22)	2759 (95% CI \pm 89; SD = 243) 49 \times (95% CI \pm 9; SD = 23)	2330 (95% CI \pm 75; SD = 206) 51 \times (95% CI \pm 8; SD = 23)	2737 (95% CI \pm 90; SD = 246) 54× (95% CI \pm 9; SD = 24)
Mean number of polymorphic RAD markers per individual	139 (95% CI \pm 18; SD = 50)	261 (95% CI \pm 24; SD = 65)	142 (95% CI \pm 21; SD = 57)	252 (95% CI \pm 26; SD = 72)
Mean number of SNPs per individual	174 (95% CI \pm 24; SD = 66)	$360 (95\% \text{ CI} \pm 24; \text{SD} = 66)$	179 (95% CI \pm 27; SD = 75)	356 (95% CI \pm 33; SD = 90)
Total Number of RAD markers in the catalog	2987	4065	2978	3925
Total number of RAD markers present in all individuals	1297	1251	1351	1426
Total number of polymorphic RAD markers present in all individuals	287	248	304	310
Number of polymorphic RAD markers with 1 SNP	220	204	232	250
RAD markers candidate neutral	164	145	167	169
RAD markers candidate balancing selection	56	59	65	79
RAD markers candidate positive selection	0	0	0	2
Number of polymorphic RAD markers with >1 SNP	29	44	72	09
RAD markers candidate neutral*	47	27	46	31
RAD markers candidate balancing selection*	9	7	8	13
RAD markers candidate positive selection*	0	0	0	1
Number of markers candidate neutral	211	172	213	200
Number of markers candidate	62	99	73	92
balancing selection				
Number of markers candidate	0	0	0	8
positive selection				
Total number of SNPs present in	365	298	388	374
all individuals				
SNPs candidate neutral	266	199	269	231
SNPs candidate balancing selection*	89	74	81	107
SNPs candidate positive selection*	0	0	0	4
Biallelic SNPs candidate neutral†	209	172	212	200

*Markers containing SNPs with conflicting classifications (e.g. one SNP candidate neutral and another candidate balancing, in the same locus) were excluded from the analyses.
†To avoid violations of the assumption of independence, only one SNP per RAD marker was used for the demographic analyses.

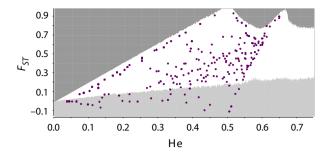


Fig. 1 Scatter plot of F_{ST} vs. expected heterozygosity (He) for the biallelic SNP loci in the analysis of genome-aligned reads without potential clones. Shaded boundaries indicate the 95% confidence intervals obtained through simulations in LOSITAN. Dark grey region indicates candidates for positive selection, and light grey region candidates for balancing selection.

selection markers identified in each analysis using genome-aligned or unaligned reads, with or without potential clones (Table 1, Fig. 1 and Fig. S4, Supporting information). Most RAD markers (~80%) contained exactly one SNP position. The majority of SNPs were biallelic (>99% overall) and none contained more than three alleles.

Approximately 40% of the candidate neutral and balancing selection markers identified in the presence of potential clones were shared between analyses with and without genome-aligned reads (Fig. 2). When the potential clones were removed, this percentage of shared markers between analyses increased slightly to 56%. Comparisons of analyses using genome-aligned reads showed that most identified markers (88% of neutral and 73% of balancing) are shared among the analyses with and without potential clones. These percentages dropped to 42% for neutral markers and 38% for balancing selection markers when using unaligned reads. Eighty-seven candidate neutral markers and 37 candidate balancing markers were common among all analyses.

Candidate loci under selection

Thirty percent (n = 37 of 124) of markers common among all four analyses, using genome-aligned or unaligned reads and with or without potential clone individuals, were inferred to be under balancing selection based on statistical comparisons. All 37 markers represented unique loci and mapped closely to single proteins in the current version of the genome of N. vectensis. Thirty-three of these were located within a gene body (19 in exons, 14 in introns, Table S1, Supporting information). The four remaining SNPs were located within a few kilobases of an annotated coding sequence. The SNP located furthest from a coding sequence was a

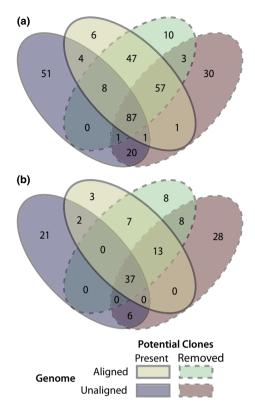


Fig. 2 Venn diagrams showing the number of markers that were unique to, and common among, the four analyses using genome-aligned or unaligned reads and with or without potential clone individuals. (a) Candidate neutral markers. (b) Candidate balancing selection markers.

polymorphism in locus number 584, which was 9067 bp from an open reading frame for a forkhead transcript factor (JGI: 239634).

All but three of the coding sequences containing or nearest to these 37 markers were annotated based on either JGI identification or through BLAST similarity. Categorization of these proteins by GO annotation suggested a diverse set of biological processes, including cellular processes, metabolic processes and response to stimulus (Fig. 3). No particular process or function (data not shown) was particularly enriched in the GO annotation; instead, these data suggest that the proteins are involved in a broad set of molecular, cellular and organismal processes. Numerically, the GO categories with the largest representation were cellular processes (n = 12, e.g. centrosomal protein), metabolic processes (n = 8, e.g. GTP-binding protein) and biological regulation (n = 8, e.g. phosphatidic acid phosphatase). KEGG annotation identified two pathways, each with one N. vectensis protein: DNA methyltransferase 1 (locus 813) involved in cysteine and methionine metabolism and natriuretic peptide receptor (locus 551) with a function in purine metabolism.

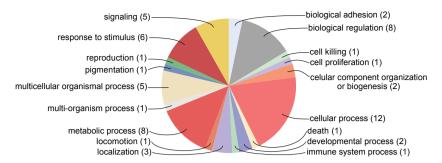


Fig. 3 Distribution of Gene Ontology (GO) categories ('biological process', level 2) for proteins coded by genes containing or most closely positioned in the genome to SNPs inferred to be under balancing selection. Analysis utilized only the 37 markers that were common among all the four analyses using genome-aligned or unaligned reads and with or without potential clone individuals. The numbers in parentheses after the GO category refer to the number of proteins annotated for each category.

For proteins near to or containing these markers under balancing selection, we identified a few proteins of particular interest due to their role in gene regulation. Two markers were located near or in a gene body for two transcription factors: locus 301 was located in an exon of the nuclear receptor co-repressor (N-CoR1) and locus 383 was 2 kb from heat shock factor 1 (HSF1). One other notable protein, a TGF β receptor (locus 729), contained one SNP located in an exon.

Demographic inferences

Principal component analyses identified three large eigenvectors (axes of variation) revealing the presence of four distinct clusters (Fig. 4, Figs S5-S7). This same result was found in all analyses using genome-aligned or unaligned reads and with or without potential clones. The eigenvector 1, with the largest eigenvalue, was not significant (P = 0.097, $\alpha = 0.05$), and the eigenvector 2, with the second largest eigenvalue, was marginally significant (P = 0.045, $\alpha = 0.05$, Table S2, Supporting information). The eigenvector 3, with the third largest eigenvalue, was highly significant (P < 0.001, $\alpha = 0.05$). All differences among identified clusters were also highly significant (Table S3, Supporting information). The results from the STRUCTURE analyses are congruent with the inferences made from the PCA (data not shown). The four identified clusters

unambiguously matched the *a priori* population assignments based on the geographic origin of the samples (Fig. 4, Figs S5–S7).

Genetic diversity in all four analyses was highest in the Massachusetts population and lowest in the Maryland population, as measured by the proportion of polymorphic markers, the expected and observed heterozygosities and the average number of alleles (Table 2). Genetic diversities for the Nova Scotia and South Carolina populations were similar, although higher in Nova Scotia. The genetic diversity estimates between analyses with or without potential clones were highly similar, but slightly higher overall when potential clones were removed.

Overall, the pairwise $F_{\rm ST}$ values suggest that differentiation was significant (i.e. 95% confidence intervals did not contain the value of the null hypothesis Fst=0) among all populations (Table 3). Genetic differentiation was greatest between the Maryland and South Carolina populations. Large genetic differentiation was also found between the Nova Scotia population and the southern populations (Maryland and South Carolina). The most similar populations were Massachusetts and Nova Scotia. None of the $F_{\rm ST}$ estimators yielded significantly different values between the analyses using genomealigned or unaligned reads. When using genomeunaligned reads, we observed significantly greater $F_{\rm ST}$ values when potential clones were included in the anal-

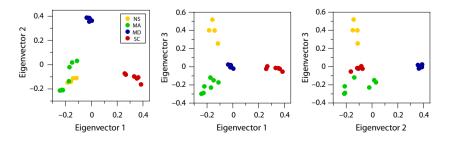


Fig. 4 Estimated population structure of *Nematostella vectensis* according to the principal component analysis (PCA) of genome-aligned reads without potential clones. Each dot represents an individual. Colours indicate the geographic site locations: Nova Scotia (NS), Massachusetts (MA), Maryland (MD) and South Carolina (SC). The three principal axes of variation are shown.

Table 2 Estimates of genetic diversity per population for four analyses (with or without clones, genome-aligned or unaligned reads)

	Nova Scotia (NS)	Massachusetts (MA)	Maryland (MD)	South Carolina (SC)
Potential clones included				
Number of samples	9	9	6	6
Genome aligned				
RAD markers with biallelic SNPs (neutral)				
Proportion of polymorphic SNP	0.517	0.612	0.239	0.445
Mean observed heterozygosity	0.213	0.290	0.112	0.161
Mean expected heterozygosity	0.169	0.205	0.082	0.143
Mean number of alleles per SNP	1.517	1.612	1.239	1.445
Number of private alleles	23	22	4	53
Unaligned				
RAD markers with biallelic SNPs (neutral)		17	2	
Proportion of polymorphic SNP	0.500	0.703	0.320	0.407
Mean observed heterozygosity	0.201	0.378	0.153	0.162
Mean expected heterozygosity	0.163	0.258	0.110	0.148
Mean number of alleles per SNP	1.500	1.703	1.320	1.407
Number of private alleles	14	12	8	25
Potential clones excluded				
Number of samples	4	7	5	6
Genome aligned				
RAD markers with biallelic SNPs (neutral)		21	2	
Proportion of polymorphic SNP	0.505	0.632	0.250	0.434
Mean observed heterozygosity	0.217	0.321	0.121	0.176
Mean expected heterozygosity	0.183	0.230	0.092	0.154
Mean number of alleles per SNP	1.505	1.632	1.250	1.434
Number of private alleles	22	28	4	40
Unaligned				
RAD markers with biallelic SNPs (neutral)		20	0	
Proportion of polymorphic SNP	0.535	0.705	0.345	0.420
Mean observed heterozygosity	0.261	0.402	0.176	0.178
Mean expected heterozygosity	0.206	0.276	0.128	0.158
Mean number of alleles per SNP	1.535	1.705	1.345	1.420
Number of private alleles	11	16	8	29

yses than when they were not. Similarly, when using genome-aligned reads, we also observed greater $F_{\rm ST}$ values when potential clones were included in the analyses than when they were not; however, these differences were not statistically significant. No significant differences were found among different $F_{\rm ST}$ estimators.

The inferred phylogeographic hypotheses clustered individuals according to their sampling location, indicating that individuals in each population share a most recent common ancestor not shared with individuals from other populations (Fig. 5). The Nova Scotia and Massachusetts populations form a monophyletic group with respect to the other two southern populations, which is consistent with the sorting of historical lineages inferred by the PCA and STRUCTURE clustering. Tree topologies of the phylogenies inferred in all analyses using genome-aligned or unaligned reads and with or without potential clones are virtually identical (Fig. 5 and Fig. S8, Supporting information), with minor differences in the bootstrap support values of the most

poorly supported branches. The number of characters used to perform the phylogenetic analyses was ~430 (Table 4). Approximately 36% of the characters were invariable in the analyses of genome-aligned reads and 47% in the analyses without genome alignment. As expected, the proportion of autapomorphic characters was greater in the analyses where potential clones were excluded.

Discussion

In this study, we have performed one of the first applications of RAD-seq to a marine invertebrate and examined genome-wide distribution of polymorphisms in natural populations of a coastal cnidarian. Together, our data reveal strong population genetic structure and clear phylogeographic relationships. Additionally, through statistical analyses of $F_{\rm ST}$ outliers, we have identified candidate regions of the genome of N. vectensis likely undergoing balancing

Table 3 Pairwise F_{ST} estimates for four analyses (with or without clones, genome-aligned or unaligned reads)

	$F_{\rm ST}^{ m W~C}$ (95% CI)	F _{ST} (95% CI)	$F_{\rm ST}^{ m Rcor}$ (95% CI)
Potential clones included			
Genome aligned			
NS vs. MA	0.298 (0.254, 0.351)	0.286 (0.239, 0.340)	0.298 (0.256, 0.350)
NS vs. MD	0.544 (0.477, 0.601)	0.556 (0.488, 0.617)	0.563 (0.498, 0.629)
NS vs. SC	0.518 (0.459, 0.572)	0.517 (0.461, 0.575)	0.521 (0.461, 0.575)
MA vs. MD	0.474 (0.426, 0.518)	0.485 (0.431, 0.536)	0.497 (0.450, 0.548)
MA vs. SC	0.480 (0.425, 0.529)	0.480 (0.429, 0.533)	0.487 (0.435, 0.538)
MD vs. SC	0.622 (0.562, 0.681)	0.617 (0.560, 0.670)	0.622 (0.564, 0.680)
Unaligned			
NS vs. MA	0.316 (0.270, 0.371)	0.303 (0.252, 0.361)	0.316 (0.268, 0.371)
NS vs. MD	0.592 (0.534, 0.648)	0.595 (0.534, 0.653)	0.602 (0.547, 0.659)
NS vs. SC	0.560 (0.498, 0.627)	0.559 (0.499, 0.618)	0.561 (0.502, 0.624)
MA vs. MD	0.460 (0.413, 0.507)	0.467 (0.417, 0.519)	0.481 (0.437, 0.533)
MA vs. SC	0.434 (0.387, 0.483)	0.438 (0.386, 0.489)	0.446 (0.397, 0.502)
MD vs. SC	0.637 (0.580, 0.697)	0.632 (0.567, 0.688)	0.637 (0.579, 0.695)
Potential clones excluded			
Genome aligned			
NS vs. MA	0.230 (0.185, 0.275)	0.218 (0.174, 0.270)	0.231 (0.183, 0.279)
NS vs. MD	0.522 (0.456, 0.590)	0.502 (0.438, 0.571)	0.508 (0.438, 0.578)
NS vs. SC	0.479 (0.422, 0.536)	0.468 (0.411, 0.527)	0.471 (0.415, 0.535)
MA vs. MD	0.431 (0.380, 0.477)	0.437 (0.389, 0.490)	0.451 (0.398, 0.503)
MA vs. SC	0.446 (0.402, 0.492)	0.440 (0.398, 0.485)	0.449 (0.408, 0.496)
MD vs. SC	0.569 (0.506, 0.636)	0.572 (0.504, 0.637)	0.576 (0.516, 0.640)
Unaligned			
NS vs. MA	0.218 (0.179, 0.258)	0.201 (0.161, 0.247)	0.221 (0.181, 0.267)
NS vs. MD	0.489 (0.430, 0.545)	0.467 (0.410, 0.525)	0.479 (0.419, 0.538)
NS vs. SC	0.494 (0.438, 0.550)	0.479 (0.431, 0.541)	0.485 (0.428, 0.538)
MA vs. MD	0.397 (0.354, 0.442)	0.395 (0.350, 0.440)	0.413 (0.368, 0.454)
MA vs. SC	0.402 (0.361, 0.446)	0.394 (0.349, 0.436)	0.406 (0.365, 0.450)
MD vs. SC	0.576 (0.520, 0.630)	0.573 (0.522, 0.627)	0.579 (0.530, 0.640)

selection in these populations. Our findings were largely insensitive to the availability of a reference genome and to the possible presence of clone individuals. These results further highlight the application of RAD-seq, and other population genomic approaches, towards understanding the genetic relationships of marine invertebrate populations and generating hypotheses about functional portions of the genome being shaped by natural selection.

RAD sequencing

Remarkably, 99% of the reads produced in this study passed as high-quality reads after our conservative filtering criteria (Fig. S1, Supporting information), indicating that RAD-seq data from cnidarians can be of extremely high quality. Given the fact that ~96% of the reads had a positive alignment to the reference genome, it can be inferred that the amount of foreign DNA contributing to the pool of RAD tags was extremely low, demonstrating that cnidarian DNA purified from nonsterile tissues (e.g. whole individuals) is suitable for

RAD sequencing. As with other population genetic studies, additional factors that could have contributed to the 4% of reads unable to produce an alignment to the reference genome include PCR errors, sequencing errors, genetic divergence and completeness of the reference genome. However, the greatest loss of data arose from reads with multiple alignments to the reference genome (24% of all reads, see Fig. S1, Supporting information). This phenomenon can be attributed to the presence of repetitive elements in the genome, which could comprise significantly large fractions of eukaryote genomes (de Koning et al. 2011), and recently duplicated genomic regions, which lack sufficient divergence for unique identification. For comparison, in the threespine stickleback study from Hohenlohe et al. (2010), 61% of the reads generated by RAD-seq (SbfI, 28-44 bp read length) produced unique alignments to the genome. Nelson et al. (2011) found that only 86% of RAD-seq reads sampled in silico from the sorghum genome (PstI and BsrFI, 36-76 bp read length) could be uniquely aligned. Thus, it is clear that there is a limit to the maximum fraction of RAD tag reads that can be

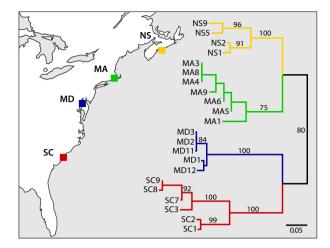


Fig. 5 Phylogeography of *Nematostella vectensis*. (left) Map showing the location of the sampled populations: Nova Scotia (NS), Massachusetts (MA), Maryland (MD) and South Carolina (SC). (right) Maximum-likelihood tree showing the most-likely phylogeographic hypothesis inferred in the analysis of genome-aligned reads without potential clones. Branches are labelled and coloured to indicate the site of collection. Numbers indicate bootstrap support values. Scale bar indicates substitutions per site.

Table 4 Statistics of the matrices used for four phylogeographic analyses (with or without clones, genome-aligned or unaligned reads)

Potential clones included	
Genome aligned	
Total number of characters	472
Proportion of invariable characters	0.364
Proportion of parsimony-informative characters	0.553
Proportion of autapomorphic characters	0.083
Proportion of missing data	0.327
Unaligned	
Total number of characters	344
Proportion of invariable characters	0.471
Proportion of parsimony-informative characters	0.439
Proportion of autapomorphic characters	0.090
Proportion of missing data	0.335
Potential clones excluded	
Genome aligned	
Total number of characters	484
Proportion of invariable characters	0.362
Proportion of parsimony-informative characters	0.519
Proportion of autapomorphic characters	0.120
Proportion of missing data	0.334
Unaligned	
Total number of characters	432
Proportion of invariable characters	0.477
Proportion of parsimony-informative characters	0.398
Proportion of autapomorphic characters	0.125
Proportion of missing data	0.408

uniquely aligned to a given reference genome. Increasing the read lengths should theoretically increase this fraction.

Based on the number of SbfI cut sites counted in the genome of N. vectensis (~2000), it was expected that ~4000 RAD markers would be obtained after sequencing. In our data set, 58% of the expected RAD markers were covered when these were identified from reads with unique alignments to the reference genome. This coverage would increase given a larger sequencing effort; plots of number of reads vs. number of RAD markers suggest that the cumulative number of covered RAD markers is close to, but has not yet reached, an asymptotic value (see Fig. S2, Supporting information). The maximum number of RAD markers that can be recovered via the analytical methods employed in this study is significantly smaller than the actual number of RAD markers present in a given genome of an N. vectensis individual. Specifically, RAD markers from repetitive regions cannot be appropriately accounted for with current methodologies. It is thus possible, if not likely, that the ratio of expected to observed number of RAD markers varies across different taxa with different genome architectures and with the kind of restriction enzyme employed. As an example, Nelson et al. (2011) achieved 57% and 73% coverage of the expected number of RAD markers in the sorghum reference genome for the BsrFI and PstI enzymes, respectively. In contrast, Hohenlohe et al. (2010) achieved ~94% coverage of RAD markers generated with SbfI in the threespine stickleback genome.

Population differentiation

The demographic inferences based on the neutral biallelic SNP markers derived from RAD loci indicated that there is strong structuring among the examined populations of N. vectensis, which span over 2000 km of coastline. Strong population differentiation and complete monophyly of populations are consistent with limited dispersal and low connectivity, as previously inferred for N. vectensis (Reitzel et al. 2008). The pairwise F_{ST} values calculated from genome-aligned markers ranged between 0.218 and 0.61. Hohenlohe et al. (2010) reported a genome-wide average F_{ST} value of 0.01 between oceanic highly-dispersing threespine stickleback populations collected 1000 km apart, whereas the values for this statistic ranged from 0.05 to 0.15 between pairs of oceanic vs. freshwater populations that have been separated for less than 10 000 years. Similarly, Roesti et al. (2012) reported F_{ST} values of 0.00–0.15 between pairs of stream vs. freshwater stickleback populations. The relatively high F_{ST} values calculated for the examined populations of N. vectensis could have been inflated if the sampled individuals were close relatives (Allendorf -Phelps effect, see Allendorf & Phelps 1981; Waples 1998). In the extreme and unlikely case that the effective number of breeders responsible for the sampled individuals in a given population (N_b) was only 2, then the maximum magnitude of the contribution of this Allendorf-Phelps effect to the observed F_{ST} values would be 0.25, calculated as $1/2(N_b)$ (Waples 1998 and citations therein). This value is smaller than most of the estimated pairwise F_{ST} values among populations of N. vectensis in this study. Another potentially important source of bias on the estimation of F_{ST} values can arise from the variable and relatively small sample sizes. The contribution of this sampling error to raw F_{ST} estimates has been shown to be ~1/(2S) (Waples 1998 and citations therein), where S is the number of individuals sampled from a population. However, the F_{ST} estimators used in this study (Weir & Cockerham 1984; Reich et al. 2009) explicitly account for this source of bias. Furthermore, a recent simulation study showed that the F_{ST}^{R} estimator is extremely accurate even when sample sizes are very small (n < 6), and that its precision is great, provided that a large number of independent markers are employed (>100) (Willing et al. 2012). Therefore, the significant, strong differentiation among populations of N. vectensis found in this study does not seem to be a methodological or analytical artefact, but is in fact a real pattern.

Possible clone individuals

The potential presence of clones among the sampled individuals had no dramatic effects on the overall population demographic inferences in this study. The main statistics that showed consistent, yet small, changes in the analyses excluding potential clones vs. analyses including potential clones were genetic diversity (Table 2) and genetic differentiation (F_{ST} values, see Table 3). Overall, the smaller genetic diversity and larger genetic differentiation observed when potential clones were included could be caused by the overrepresentation of particular genotypes and biases in the allelic differences among populations. The decrease in population sample sizes after the exclusion of potential clones could have also magnified the effect of sampling error and thus contributed to the observed small changes in the values of these statistics.

Phylogeography

We found a clear genetic break between northern (NS, MA) and southern (MD, SC) populations, but a colonization scenario that could explain this pattern is unclear. Principally, we were unable to root the tree due to the uncertainty in the history of these populations and the lack of a clear outgroup species. Because the northern portion of the range of *N. vectensis* was

covered during the last glacial maximum, a reasonable hypothesis would be that populations recolonized estuaries north of Cape Cod after the glaciers receded, similar to other coastal invertebrates (Jennings et al. 2009). Thus, we would expect reduced genetic diversity in these higher-latitude populations. However, genetic diversity was overall higher in these more northern populations. Similarly, genetic diversity assayed with AFLPs (Reitzel et al. 2008) and by sequence-based markers (Reitzel et al. 2008; Sullivan et al. 2009; Reitzel et al. 2010) also suggested that genetic diversity is similar or even higher in populations north of Cape Cod. Previous research with the estuarine fishes Fundulus heteroclitus (Adams et al. 2006; Williams & Oleksiak 2008) and Menidia menidia (Mach et al. 2011), both of which have overlapping ranges with N. vectensis, has also observed similar genetic diversity among populations along the Atlantic coast of North America. In these fish species, the absence of reduced diversity in higher-latitude populations is in part a result of local adaptation along environmental clines and in response to anthropogenic stressors. These environmental variables have shaped the regional genetic diversity despite the movement of populations during glacial periods. Future research with N. vectensis incorporating additional locations along the Atlantic coast of North America may help resolve the directionality of population colonization and the importance of genetic adaptation to regional environmental conditions.

Utility and promise for nonmodel organisms

Local physical oceanographic processes and humanmediated introductions can greatly influence the population connectivity dynamics among estuarine communities. The life history of N. vectensis, containing an egg mass that retains embryos, a demersal larva with a short swimming period (<7 days) and an infaunal adult, would likely promote limited dispersal of adults and developmental stages. Consistent with this expectation, surveys of genetic structure within estuaries and between adjacent locations have identified significant structure (Reitzel et al. 2008). Previous genetic research has indicated that anthropogenic dispersal has played an important role in shaping the broad geographic scale distribution and resulting population genetic relationships in N. vectensis (Darling et al. 2004; Reitzel et al. 2008; Darling et al. 2009). Similar to a number of other coastal invertebrates in North America, N. vectensis appears to have been introduced from the Atlantic coast to the Pacific coast, potentially through the transport of commercial shellfish. The addition of the high-density SNP data generated in this study to previous data will provide a high degree of analytical

power to understand both genetic partitioning in the small spatial scales of natural dispersal and large scales of long-distance anthropogenic dispersal. Even more so, these methods hold great opportunity for understanding similar processes in other coastal species. Despite the differences in the number of loci and SNP recovered when reads were filtered with the genome and when they were not, the results from the demographic inferences were overall identical. Furthermore, the use of the reference genome did not substantially affect the number of retrieved RAD loci across populations, thus avoiding the introduction of an ascertainment bias. Our results highlight the usefulness of RAD sequencing for population genetics and evolutionary studies with or without the availability of a reference genome (for more examples, see Baird et al. 2008; Emerson et al. 2010; Amores et al. 2011; Baxter et al. 2011; Dasmahapatra et al. 2012; Peterson et al. 2012). Because most coastal and oceanic species from shallow and deep environments lack genomic resources, RAD-seq offers a valuable tool for the identification of native source locations for introduced species, and a tremendous opportunity for the characterization of genetic diversity in other species of ecological or conservation interest, especially those for which basic taxonomic and population structure knowledge has been particularly challenging to obtain (e.g. octocorals, see Herrera et al. 2010; McFadden et al. 2010 and references therein; Herrera et al. 2012).

Selection

High-density SNP maps generated from field-sampled populations can be used to identify genomic regions potentially under selection. When correlated with known phenotypic diversity, linkage studies provide a powerful tool in functional genomics to bridge genetic and phenotypic variation (Feder & Mitchell-Olds 2003; Mitchell-Olds et al. 2008; Stinchcombe & Hoekstra 2008; Nadeau & Jiggins 2010). RAD-seq and similar methods, for example restriction-site tiling analysis (Pespeni et al. 2010), that generate large number of SNPs provide the technological approaches to produce these data for nonmodel species. For example, studies in stickleback (Hohenlohe et al. 2010) and the purple sea urchin (Pespeni et al. 2012) have each identified novel genomic regions under selection, which correlate with differential phenotypes in natural populations. Given the extensive latitudinal range and high degree of genetic structure of N. vectensis, it is reasonable to expect local adaptation in its populations.

Two previous studies with *N. vectensis* mined SNPs from Sanger-sequenced expressed sequence tags and identified geographically segregated polymorphisms in

highly conserved regions of genes (Reitzel et al. 2010), one of which has dramatic functional impacts on protein function (NF-kB, Sullivan et al. 2009). This previous approach has clear limitations because SNPs could only be identified in coding regions, which are certainly important in adaptive evolution (Hoekstra & Coyne 2007), but would not identify SNPs in noncoding regions that are also of functional importance (Wray 2007). Furthermore, this approach introduces biases, such as ascertainment bias, because all source sequences for SNP identification are generated from individuals collected at one geographic location. In this study, we have identified SNPs throughout the genomes of individuals collected from four geographic locations. We utilized the restriction enzyme SbfI to generate the RAD tags, which would at most produce ~2000 cuts based on counts from the reference genome of N. vectensis. This number is considerably smaller than the number of cut sites in the genomes of teleost fishes (~25 000-30 000), such as the threespine stickleback (Hohenlohe et al. 2010; Amores et al. 2011), which makes it impossible to generate equivalent high-density mapping for N. vectensis from data generated with this same restriction enzyme. To achieve a higher mapping density, additional, more frequently cutting restriction enzymes would be required (e.g. EcoRI). However, even under this restriction, we identified 37 polymorphic sites common among all analyses that were inferred to be under balancing selection. Perhaps surprisingly, a large majority of these SNPs were in gene bodies, many of which have clear orthology to proteins of known function in other animals. For example, one SNP was located in an intron of a single ortholog to DNA methyltransferase 1, an enzyme that establishes and regulates tissue-specific patterns of cytosine methylation, and an intergenic SNP was located nearest to heat shock factor I, the principle transcription factor that regulates downstream expression of genes involved in temperature stress. Future research utilizing a more frequently cutting enzyme will generate a higher-density SNP map, which will facilitate a more thorough analysis of genomic regions under selection in these populations.

Future directions for Nematostella

In addition to resolving population relationships and identification of genomic regions undergoing selection, RAD-seq identification of SNPs can be used as a tool to push functional molecular studies in *N. vectensis*. Identification of SNPs linked to a particular genomic region will allow researchers to identify and test the relationship of candidate genomic loci with phenotypes of interest. *N. vectensis* has emerged as a premier model

in cnidarian developmental biology and is a prime candidate as an experimental system in functional molecular genetics. Experimentally induced mutations combined with SNP profiling are a powerful tool that can be used to identify mutations underlying novel phenotypes in N. vectensis with high resolution. Researchers would be able to exploit the asexual reproductive biology of N. vectensis to perpetually maintain deleterious alleles in heterozygous individuals, which would facilitate conducting forward genetic screens to investigate molecular mechanisms governing the development of particular morphological characters or differences in physiology. This unbiased forward approach would be an influential technological leap for evolutionary developmental biology and evolutionary ecology of cnidarians, which, until now, rely heavily on candidate gene approaches. Such unbiased approaches would inherently investigate novel mechanisms governing biological processes.

Conclusion

We have presented the broad utility of RAD-seq to characterize the genome-wide distribution of polymorphisms in a coastal invertebrate. Our data reveal strong population genetic structure, clear phylogeographic relationships and candidate regions of the genome undergoing selection in natural populations. This approach holds tremendous promise for understanding the genetic relationships and phylogeography of other marine invertebrates, including those of conservation concern that have traditionally been difficult to study due to the lack of genetic variation (e.g. corals). Population genomic approaches will produce data that may be used to quantify the role played by the environment in selecting for local adaptation via ecologically important regions of the genome. This in turn will generate hypotheses about how functional portions of the genome are being shaped by natural and anthropogenic selection.

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A.M.R., S.H. and M.J.L. designed the experiment. A.M.R. and S.H. analysed the data and drafted the manuscript. M.Q.M. and T.M.S. participated in design of the study and interpretation of results. All authors approved the final manuscript.

Data accessibility

- Raw DNA sequence reads are available at the U.S.
 National Center for Biotechnology Information
 (NCBI) Sequence Read Archive (SRA) accession
 number SRA055050.
- Alignments of RAD sequence reads to the reference genome, as produced by BOWTIE, and the RAD markers used in the population genomic analyses are available in DRYAD doi:10.5061/dryad.gk2vc.

 Genomic positions, protein IDs, population statistics and GO results for each one of the candidate loci under selection are available as supplementary material.

Supporting information

Additional supporting information may be found in the online version of this article.

Fig. S1 Percentages of the number of reads retained after each major filtering step.

Fig. S2 Scatter plots of the coverage per locus and the number of identified RAD markers a functions of the number of reads generated per individual.

Fig. S3 Box and whisker plots showing the effect of utilizing the *N. vectensis* reference genome to filter sequence reads. Data are from resampled data sets and are grouped by population.

Fig. S4 Scatter plots of $F_{\rm ST}$ vs. expected heterozygosity (He) for the biallelic SNP in the four analyses.

Fig. S5 PCA eigenvector 1 vs. eigenvector 2, results of the four analyses.

Fig. S6 PCA eigenvector 1 vs. eigenvector 3, results of the four analyses.

Fig. S7 PCA eigenvector 2 vs. eigenvector 3, results of the four analyses.

Fig. S8 Phylogeography of *N. vectensis*, results of the four analyses. Presented trees show the most-likely phylogeographic hypotheses resulting from maximum-likelihood analyses.

Table S1. Details of the 37 markers candidate under balancing selection common among all four analyses.

Table S2. Tracy-Widom statistics of significance of the identified principal components.

Table S3. Chi-square statistics of significance of the differences between identified populations identified by the PCA.